



Test-day milk yield as a selection criterion for dairy buffaloes (*Bubalus bubalis* Artiodactyla, Bovidae)

Humberto Tonhati¹, Mário Fernando Cerón-Muñoz², João Ademir de Oliveira¹, Lenira El Faro³,
André Luís Ferreira Lima¹ and Lucia Galvão de Albuquerque¹

¹Faculdade de Ciências Agrárias e Veterinárias, Universidade Estadual Paulista “Júlio de Mesquita Filho”,
Jaboticabal, SP, Brazil.

²Universidad de Antioquia, Medellín, Colombia.

³Agência Paulista de Tecnologia dos Agronegócios, Ribeirão Preto, SP, Brazil.

Abstract

Due to the great demand for buffalo milk by-products the interest in technical-scientific information about this species is increasing. Our objective was to propose selection criteria for milk yield in buffaloes based on total milk yield, 305-day milk yield (M305), and test-day milk yield. A total of 3,888 lactations from 1,630 Murrah (*Bubalus bubalis*) cows recorded between 1987 and 2001, from 10 herds in the State of São Paulo, Brazil, were analyzed. Covariance components were obtained using the restricted maximum likelihood method applied to a bivariate animal model. Additive genetic and permanent environmental effects were considered as random, and contemporary group and lactation order as fixed effects. The heritability estimates were 0.22 for total milk yield and 0.19 for M305. For test-day yields, the heritability estimates ranged from 0.12 to 0.30, with the highest values being observed up to the third test month, followed by a decline until the end of lactation. The present results show that test-day milk yield, mainly during the first six months of lactation, could be adopted as a selection criterion to increase total milk yield.

Key words: genetic correlations, genetic parameters, heritability, selection efficiency.

Received: July 20, 2007; Accepted: February 25, 2008.

Introduction

Milk production in buffaloes (*Bubalus bubalis* Artiodactyla, Bovidae) is an economically important trait, accounting for more than 40% of the milk produced in Asia and sustaining a powerful dairy industry in Italy, with buffalo milk also being consumed on a smaller scale in many other countries.

Asian countries traditionally tend to express milk production based on yields at 305 days of lactation (Sane *et al.*, 1972; Mourad and Mohamed, 1995), while in Italy, taking into account buffaloes mean lactation period, milk yield adjusted to 270 days of lactation is adopted as a selection criterion (Rosati and Van Vleck, 2002) as recommended by the International Committee of Animal Recording (ICAR, 2008). In dairy cattle, 305 days cumulative milk yields or yields at partial periods of lactation have been considered for genetic evaluation. For the calculation of these yields,

milk production is generally monitored at average intervals of 30 days and the production of one cow on each test day is called test-day (TD) milk yield. Cumulative yields for different lactation periods are calculated using estimated extension or adjustment factors, with accuracy of cumulative yield depending on the quality and quantity of the available test-day records (Tonhati *et al.*, 2004; ICAR, 2008).

There are various methods to calculate the milk production in the lactation (ICAR, 2008), but the usual one is the test interval method as described in Everett and Carter (1968). The problem of using estimates of extension factors is that this approach assumes that there is no variability in the shape of the lactation curve between animals, thus eliminating some genetic variation in milk yield (Shahrbabak MM – PhD Thesis, University of Guelph, Guelph, Canada, 1997). On the other hand, exclusion of short or incomplete lactation records might cause bias due to the pre-selection of data. In addition, the use of total yields or yields accumulated in long lactation periods implies not only the impossibility of considering environmental effects that may influence milk production during certain periods but also of including in the analysis incomplete data from young animals.

Send correspondence to Humberto Tonhati. Departamento de Zootecnia, Faculdade de Ciências Agrárias e Veterinárias, Universidade Estadual Paulista “Júlio de Mesquita Filho”, Via de Acesso Prof. Paulo Donato Castellane s/n, Prédio II, 14884-900 Jaboticabal, SP, Brazil. E-mail: tonhati@fcav.unesp.br.

An alternative to overcome the problem of extending production records until the end of lactation using extension factors is to utilize TD yields instead. The main advantage of this approach is the possibility of reducing the number of TD, thus reducing the generation interval. In addition, a cow can be included in the evaluation if it has at least one measurement and the larger amount of available data will increase sires evaluation accuracy and minimize possible bias due to excluding incomplete lactations (Schaeffer *et al.*, 1977; Ptak and Schaeffer, 1993).

In dairy cattle, heritability estimates for TD yields have been close to or slightly less than those obtained for 305-day yields (Meyer *et al.*, 1989), however, applying random regression models heritability estimates for daily yields have been higher than for 305-d yields (Jamrozik and Schaeffer, 1997). In general, rank correlations between animals' breeding values for 305-day milk yields and some TD milk yields are relatively high, with a large number of sires in common when sires are ranked based on the two criteria (Ptak and Schaeffer, 1993; Swalve, 1995).

In buffaloes, (Tiwana MS and Dhillon JS – Proceedings of I World Buffalo Congress, Cairo, Egito, 1985) a heritability estimate of 0.37 for cumulative milk yield up to the fourth month of lactation and a genetic correlation of 0.96 between this trait and 305-day milk yield have been reported. The authors concluded that the adoption of partial production as a selection criterion may result in greater annual genetic gain, mainly due to reduction in the generation interval.

Based on the above considerations, the objectives of the present study were to estimate genetic parameters for total, 305-day and TD milk yields and to propose a selection criterion for milk production in buffaloes.

Materials and Methods

Data

The data analyzed were obtained from the buffalo milk-recording program maintained by the Department of Animal Science, UNESP, Jaboticabal, São Paulo, Brazil, including 10 farms in the State of São Paulo. Most of the data (32%) originated from one farm only, which started a recording system in 1987. There were a minimum of 42 and a maximum of 1,225 lactations per herd. Data included 3,888 lactations (from first to sixth order) from 1,630 Murrah breed cows (*Bubalus bubalis*) recorded from 1987 through 2001. The number of lactations per year varied from 62 to 721. Animals were reared on pastures mainly of *Brachiaria* and *Panicum*. In general, the practice of bulk feed supplementation based on chopped sugar cane or grass silage and cotton seed-, barley- and soybean-based concentrate was adopted mainly during the dry season (April to September). Mineral salt was offered regularly and a system of natural breeding and artificial insemination was adopted. Sanitary control was performed according to the

recommendations of the Brazilian Ministry of Agriculture (<http://www.agricultura.gov.br/>). Drugs against ecto- and endoparasites were applied twice a year on average. Test-day records were obtained monthly. The cows were milked twice a day and the calves were with the cow during milking.

Total milk yield was defined as the amount of milk (kg) produced throughout the lactation period, while M305 was obtained by truncating the production on day 305 of lactation. These traits were chosen because Madalena *et al.* (1992) found in Zebu cattle that adjusting lactation milk yield for days in milk removed genetic variability of milk yield. Moreover, in previous work, with Buffaloes, Tonhati *et al.* (2004) reported less genetic variability for lactation milk yield when adjustment factors for days in milk were used. Average lactation length for this population was 254.42 days with a standard deviation of 66.11 days. Test-day yields refer to the monthly milk measurements performed throughout the lactation period (TD1 to TD10).

Model

Test-day milk yield, total milk yield and M305 (co)variance components were estimated by restricted maximum likelihood (REML) using a bivariate animal model according to the algorithm developed by Boldman *et al.* (1995). The model can be represented as:

$$y_i = X_i b_i + Z_i a_i + W_i c_i + e_i;$$

where y_i is the observation vector for the i th trait (test-day milk yield, total milk yield or M305), and X_i , Z_i , and W_i are incidence matrices related to fixed (b_i) and random animal (a_i) and permanent environmental (c_i) effects, respectively.

It was assumed that:

$$E[y] = [X]; E \begin{bmatrix} a_1 \\ a_2 \\ c_1 \\ c_2 \\ e_1 \\ e_2 \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \\ 0 \\ 0 \\ 0 \\ 0 \end{bmatrix};$$

$$\text{Var} = \begin{bmatrix} a_1 \\ a_2 \\ c_1 \\ c_2 \\ e_1 \\ e_2 \end{bmatrix} = \begin{bmatrix} \sigma_{a_1}^2 A & \sigma_{a_1 a_2} A & 0 & 0 & 0 & 0 \\ \sigma_{a_2 a_1} A & \sigma_{a_2}^2 A & 0 & 0 & 0 & 0 \\ 0 & 0 & \sigma_{c_1}^2 I & \sigma_{c_1 c_2} I & 0 & 0 \\ 0 & \sigma_{c_2 c_1} I & \sigma_{c_2}^2 I & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & \sigma_{e_1}^2 I & \sigma_{e_1 e_2} I \\ 0 & 0 & 0 & 0 & \sigma_{e_2 e_1} I & \sigma_{e_2}^2 I \end{bmatrix}$$

where A is the relationship matrix, $\sigma_{a_i}^2$ are additive genetic variances, σ_{a_i, a_j} are additive genetic covariances, $\sigma_{c_i}^2$ are permanent environmental variances, σ_{c_i, c_j} are permanent environmental covariances, $\sigma_{e_i}^2$ are residual variances and σ_{e_i, e_j} are residual covariances between two traits i and j and I = identity matrix of appropriate order.

The model included contemporary group (herd-calving year for total milk yield or 305-d milk yield and herd-date of recording for test-day milk yield) and lactation order as fixed effects. Calving season was not included as a fixed effect since previous analyses had shown no effect of this factor on milk yield (Tonhati *et al.*, 2000a). Six lactation orders and 67 contemporary groups with a minimum of five lactation records were considered for analysis. There were 7,396 animals in the relationship matrix from which 2,062 (27.88%) were inbred with a mean inbreeding coefficient of 4.8%. Results presented for test day milk yields (TD1 to TD10) were obtained from bivariate analysis with total milk yield.

Expected genetic gain and correlated response to selection were obtained using the estimates of heritability (h^2), genetic correlations and phenotypic standard deviations. Selection of the best 5% males was examined, corresponding to a selection intensity factor of 2.06 (Lush, 1964). Females were randomly replaced in the herd, with selection intensity equal to zero. Thus, the mean selection intensity factor for all traits was 1.03. Expected direct and correlated responses to selection and the relative efficiency of response were calculated by the usual selection index formulas, considering a progeny test with number of daughters by sire varying from 5 to 100.

Results and Discussion

The observed mean for total milk yield (kg) was $1,532.68 \pm 660.88$ and $1,495.08 \pm 617.12$ for M305, which is similar to those reported by Tonhati *et al.* (2000a). The number of observations and the means, standard deviations and coefficients of variation for TD milk yield are shown in Table 1.

Monthly TD milk yield means varied according to the phase of lactation, with the highest production being recorded for the second month of lactation, followed by a gradual decrease through to the tenth month. The increase in milk yield was 7.80% from the first to the second month, corresponding to a milk yield of 0.55 kg. The mean percentage of decline observed after the second test-day was 7.09%. A minimal decline in percentage of production was observed from the second to the third month (4.48%), while the largest decrease (9.82%) occurred between the eighth and ninth month of test. These results support the recommendation from ICAR (2008) for using milk yield adjusted for 270 days of lactation.

In absolute terms, the average monthly decline in milk yield was 0.42 kg from the lactation peak. The largest decrease (0.52 kg) was observed from the fourth to the fifth test-day, while the smallest decrease (0.30 kg) occurred between the ninth and tenth test-day. The percentage and absolute total losses in milk yield from the second to the tenth test-day were 44.46% and 3.39 kg, respectively. Similarly, Bremner-Gaona *et al.* (1985) for Mediterranean and Mu-

Table 1 - Number of observations (N) and mean, standard deviation (SD), coefficient of variation (CV) and minimum and maximum values of test-day milk yield in kg (TD1 to TD10).

Variable	N	Mean (kg)	SD (kg)	CV (%)	Minimum (kg)	Maximum (kg)
TD1	3078	7.04	2.64	37.45	1.30	20.50
TD2	3075	7.59	2.85	37.54	1.00	20.90
TD3	3047	7.25	2.84	39.14	1.10	23.50
TD4	2959	6.81	2.68	39.42	1.00	22.00
TD5	2865	6.29	2.55	40.57	1.00	19.00
TD6	2613	5.82	2.41	41.34	1.00	18.80
TD7	2460	5.41	2.18	40.34	0.70	15.50
TD8	2053	4.99	2.02	40.58	0.80	16.80
TD9	1541	4.50	1.87	41.54	0.50	15.60
TD10	960	4.20	1.73	41.21	0.50	15.40

nõz-Berrocal *et al.* (2005) for Murrah buffalo in Brazil reported that lactation peaked at the second month of lactation, with milk yield decreasing continuously thereafter. In Brazil, until few years ago, selection for milk yield in buffaloes was based on the milk production of dams. Tonhati *et al.* (2000b) reported milk yield average breeding values for Murrah buffalo sires close to zero from 1977 to 1995, showing that the selection was not effective.

Milk yields on each test-day had high coefficients of variation (Table 1), indicating considerable variability in the shape of the lactation curve. The lower coefficients of variation obtained for test-days at the beginning of lactation might be due to the greater standardization of post-partum treatment and/or to the body status of the cows which generally have good body reserves after calving. On the other hand, the smaller number of observations close to the end of lactation, as well as individual differences between animals in terms of lactation persistency, might have contributed to the higher coefficients of variation for those test-days. Considering the positive genetic correlation between persistency of lactation and total milk yield (Gengler, 1996), it could be expected that only the more productive cows would be lactating in the last stage of lactation.

Table 2 shows the estimates of variance components and of genetic and phenotypic parameters obtained by bivariate analysis including total milk yields and M305. The heritability estimates were within the limits reported in the literature for these traits (Tonhati *et al.*, 2000a; Rosati and Van Vleck, 2002). A high positive genetic correlation (0.99) was also observed between total milk yield and M305, demonstrating a pleiotropic effect of genes on these traits.

Estimates of permanent environmental correlations were high, suggesting that both traits are affected in part by the same environmental factors. Consequently, the repeatability estimates indicate that, as an aid for selection, past

Table 2 - Estimates of additive genetic (σ_a^2), permanent environmental (σ_c^2) and residual (σ_e^2) variances (kg^2), and estimates of heritability (h^2), repeatability (t), genetic (r_a), permanent (r_c) and temporary environmental (r_e) correlations for total milk yield (TY) and yields at 305 days of lactation (M305).

Parameter	Total milk yield	M305
σ_a^2	56035.86	41057.26
σ_c^2	54081.86	49063.31
σ_e^2	150007.01	121019.82
h^2	0.22	0.19
t	0.42	0.43
$r_a = 0.99$	$r_c = 1.00$	$r_e = 0.93$

values for both, total milk yield and M305, can be used as indicators of future yields. Estimates of variance components and of genetic and phenotypic parameters for test-day milk yields obtained from bivariate analyses, with total milk yield, are shown in Table 3.

Phenotypic variance estimates showed the same trend as test-day milk yield, increasing from the first to the second test-day and decreasing thereafter (Figure 1). The highest estimates of residual and additive genetic variances were observed in the first and second months of lactation, respectively, followed by a gradual decrease during the subsequent periods of lactation (Table 3). Estimates of permanent environmental variance showed an increase from the beginning of lactation to the third test-day (TD3), declining thereafter. This variance behavior led to higher heritability estimates during the initial phase of lactation, with a maximum value of 0.30 at the third test-day (Figure 1). Although all the variance estimates decreased with days in milk, the decrease in the genetic variance was more than in the residual variances and, as a consequence, heritability estimates were lower after the sixth month of lactation. Probably, this pattern is due to the fact that only more persistent cows remain producing until the tenth month of lactation which would decrease genetic variability. Moreover, higher heritability estimates during the first half of lactation might be due to better control of environmental conditions through standardization of post-partum management. Estimates of heritability for TD yields reported in the literature for dairy cattle show different magnitudes but, in general, the mid-lactation yields are more heritable than those at the beginning and at the end of lactation (Meyer *et al.*, 1989; Pandner *et al.*, 1992 and Swalve, 1995).

The advantage of using TD yields for genetic evaluations due to their higher coefficients of heritability (Meyer *et al.*, 1989; Ptak and Schaeffer, 1993; Swalve, 1995) was confirmed in the present study. In addition, the genetic correlations between TD yields and total milk yield were high (Table 3), indicating that a large proportion of the additive genetic variance is common to both TD and total yields.

Table 4 shows estimates of genetic gain (ΔG) for total milk yield (TY), correlated responses and relative selection

Table 3 - Estimates of additive genetic (σ_a^2), permanent environmental (σ_c^2) and residual (σ_e^2) variances (kg^2), and estimates of heritability (h^2), repeatability (t), genetic (r_a), permanent (r_c) and temporary environmental (r_e) correlations obtained for test-day milk yield (TD1 to TD10) by bivariate analysis with total milk yield.

Variable	σ_a^2	σ_c^2	σ_e^2	r_a	r_c	r_e	h^2	t
TD1	0.89	0.44	2.50	0.94	1.00	0.55	0.23	0.35
TD2	1.27	0.69	2.48	0.98	0.93	0.74	0.29	0.44
TD3	1.25	0.78	2.16	0.99	0.98	0.74	0.30	0.48
TD4	0.98	0.56	2.19	1.00	1.00	0.74	0.26	0.41
TD5	0.87	0.54	2.05	0.96	1.00	0.75	0.25	0.41
TD6	0.75	0.48	1.81	1.00	1.00	0.73	0.25	0.40
TD7	0.51	0.59	1.60	0.95	0.96	0.70	0.19	0.41
TD8	0.41	0.34	1.70	0.93	0.93	0.64	0.17	0.31
TD9	0.26	0.44	1.53	0.96	0.91	0.59	0.12	0.32
TD10	0.21	0.18	1.24	0.82	0.52	0.45	0.13	0.24

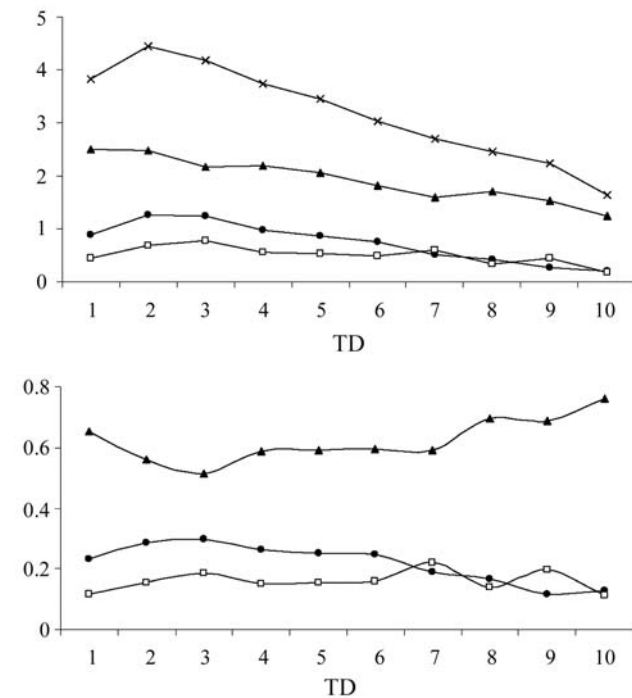


Figure 1 - Estimates of phenotypic (x), genetic (●), permanent environmental (□) and residual (▲) variances (top) and estimates of heritability (●) and permanent environmental (□) and residual (▲) variances as proportions of phenotypic variance (bottom) for test-day milk yield (TD).

efficiencies for total milk yield as a function of direct selection for M305 and TD yields. The estimates of correlated response for total milk yield as a function of direct selection for TD yields indicate that the adoption of initial TD yields as a selection criterion, *i.e.*, from second to sixth TD, would be an efficient approach to increase total milk yield, because these TD milk yields have higher heritability estimates than TY and the genetic correlations, among the TD

Table 4 - Expected direct response to selection for total milk yield (TY) and correlated response for total milk yield with direct selection for test-day milk yield (TD) and relative selection efficiency.

Trait	Correlated response (number of daughters by sire)				Relative selection efficiency (number of daughters by sire)			
	5	10	50	100	5	10	50	100
TY*	115.8	147.8	210.3	225.2	100.0	100.0	100.0	100.0
M305	107.8	139.2	203.9	220.3	93.1	94.2	97.0	97.8
TD1	110.8	141.0	198.8	212.4	95.7	95.3	94.5	94.3
TD2	126.6	158.2	213.2	224.9	109.4	107.0	101.3	99.8
TD3	129.6	161.5	216.1	227.7	111.9	109.2	102.7	101.1
TD4	123.8	156.1	214.8	227.9	106.9	105.5	102.1	101.2
TD5	117.0	148.0	205.2	218.2	101.1	100.1	97.5	96.9
TD6	121.9	154.2	213.8	227.3	105.3	104.2	101.6	100.9
TD7	103.5	133.6	195.6	211.4	89.3	90.3	93.0	93.8
TD8	96.6	125.7	188.2	204.8	83.4	85.0	89.5	90.9
TD9	85.7	113.7	182.4	203.4	73.9	76.9	86.7	90.3
TD10	75.9	100.2	158.2	175.5	65.4	67.7	75.2	77.9

*Expected direct genetic gain.

and TY, are greater than 0.95. However, this advantage would decrease with an increase in the number of daughters by sire because, in this case, differences in accuracy by using TD or TY as selection criterion are small. For example, for the 5th test-day milk yield, if the sire has 5 daughters, the accuracy of selection would be 0.50 for TD5 and 0.47 for TY. However, with 100 daughters the accuracy would be 0.93 for TD5 and 0.92 for TY. But, in some cases, to obtain a larger number of daughters per sire it would be necessary to use the sires for a longer period, thus increasing the generation interval and decreasing response to selection. The genetic parameters estimated indicate that milk yields around the lactation peak, *i.e.*, the second and third TD, have the highest heritability and a strong relationship with total milk yield and, therefore, result in the greatest expected correlated responses for TY. Based on these results, an economic selection index considering two or more TD, including other measurements such as persistency of lactation, could be developed. To favor animals with higher persistency of lactation, TD after the peak of production probably should be included in the selection index.

The results obtained in the present study indicate that adoption of test-day milk yields as selection criteria might contribute to greater genetic gain in total milk production. In addition, the results indicate that test-day milk yields near the lactation peak (TD2 and TD3) are most closely related to total milk yield.

Acknowledgments

Appreciation is extended to the Fundação de Apoio à Pesquisa do Estado de São Paulo (FAPESP) and Conselho

Nacional de Desenvolvimento Científico e Tecnológico (CNPq) for financial support.

References

- Boldman KG, Kriese LA, Van Vleck LD and Kachman SD (1995) A Manual for Use of MTDFREML. Department of Agriculture/Agricultural Research Service, Lincoln, 120 pp.
- Bremner-Gaona JH, Andrade VJ, Pereira CS and Sampaio IBM (1985) Estudo da curva de lactação em búfalas da raça Mediterrânea e seus mestiços. Arq Bras Med Vet Zootec 37:477-495 (Abstract in English).
- Everett RW and Carter HW (1968) Accuracy of test interval method of calculating dairy herd improvement association records. J Dairy Sci 51:1936-1941.
- Gengler N (1996). Persistency of lactation yields: A review. Interbull Bull 12:87-96.
- Jamrozik J and Schaeffer LR (1997) Estimates of genetic parameters for a test day model with random regressions for yield traits of first lactation Holsteins. J Dairy Sci 80:762-770.
- Lush JL (1964) Como a seleção muda uma população - Expressão dos resultados. In: Melhoramento Genético dos Animais Domésticos. USAID, Rio de Janeiro, pp 186-217.
- Madalena FE, Lemos AM and Teodoro RL (1992) Consequences of removing the variation in lactation length on the evaluation of dairy cattle breeds and crosses. Braz J Genet 15:585-594.
- Meyer K, Graser HU and Hammond K (1989) Estimates of genetic parameters for first lactation test day production of Australian Black and White cows. Livestock Prod Sci 21:77-199.
- Mourad KA and Mohamed MM (1995) Genetic and phenotypic aspects of milk yield traits and reproductive performance of Egyptian buffaloes. Egypt J Anim Prod 32:125-137.
- Munõz-Berrocal M, Tonhati H, Cerón-Muñoz M, Duarte JMC and Chabariberi RL (2005) Uso de modelos lineares e não lineares para o estudo da curva de lactação em búfalos Murrah e seus mestiços em sistema de criação semi extensivo no Estado de São Paulo. Arch Latinoam Prod Anim 13:19-23 (Abstract in English).
- Pander BL, Hill WG and Thompson R (1992) Genetic parameters of test day records of British Holstein-Friesian heifers. Anim Prod 55:11-21.
- Ptak E and Schaeffer LR (1993) Use of test day yields for genetic evaluation of dairy sires and cows. Livestock Prod Sci 34:23-34.
- Rosati A and Van Vleck LD (2002) Estimation of genetic parameters for milk, fat, protein and mozzarella cheese production in the Italian river buffalo population. Livestock Prod Sci 74:185-190.
- Sane DD, Khanna RS, Bajpai LD and Bhat PN (1972) Studies on Murrah buffalo (*Bubalus bubalis*). II: Genetic analysis of milk yield and peak yield. Indian J Anim Prod 3:61-65.
- Schaeffer LR, Minder CE, McMillan I and Burnside EB (1977) Non-linear techniques for predicting 305-day lactation production of Holstein and Jerseys. J Dairy Sci 60:1636-1644.
- Swalve HH (1995) The effect of test day models on the estimation of genetic parameters and breeding values for dairy yield traits. J Dairy Sci 78:929-938.
- Tonhati H, Muñoz MFC, Duarte JMC, Reichert RH, Oliveira JA and Lima ALF (2004) Estimates of correction factors for

lactation length and genetic parameters for milk yield in buffaloes. *Arq Bras Med Vet Zootec* 56:251-257 (Abstract in English).

Tonhati H, Muñoz MFC, Oliveira JA, Duarte JMC, Furtado TP and Tseimazides SP (2000a) Parâmetros genéticos para a produção de leite, gordura e proteína em bubalinas. *Rev Bras Zootec* 29:2051-2056 (Abstract in English).

Tonhati H, Vasconcelos FB and Albuquerque LG (2000b) Genetic aspects of productive and reproductive traits in a Murrah buffalo herd in São Paulo, Brazil. *J Anim Breed Genet* 117:331-336.

Internet Resource

ICAR (2008) Guidelines for buffalo milk recording for low to medium and medium to high input production systems. http://www.icar.org/pages/recording_guidelines.htm (February 13, 2008).

Associate Editor: Pedro Franklin Barbosa

License information: This is an open-access article distributed under the terms of the Creative Commons Attribution License, which permits unrestricted use, distribution, and reproduction in any medium, provided the original work is properly cited.